# Random signals 

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## Random signals

- deterministic signals (can be represented by an equation) have one substantial drawback - they carry very little information (for example cosine: amplitude, frequency, phase shift).
- real-world signals are tough to be described as deterministic (for example physical model of speech production is very complex and anyway simplified).
$\Rightarrow$ in signal theory, we will consider these useful signals as random signals (processes) (for example speech, circulation of letters in agencies of Czech Posts, exchange rate CZK/EUR ...).

According to the character of the time axis, random signals are divided into continuous time random signals (the time is defined for all $t$ ) and discrete time random signals (only for discrete $n$ ).

These signals can not be represented in all time-points (in this case, they would be deterministic), we will rather look for characteristic properties of random signals such as mean value, probability density function, etc.

## Definition of random process

- continuous time: the system $\left\{\xi_{t}\right\}$ of random variables defined for all $t \in \Re$ is called random process, it is denoted $\xi(t)$.
- discrete time: the system $\left\{\xi_{n}\right\}$ of random variables defined for all $n \in N$ is called random process, it is denoted $\xi[n]$.


## Set of realizations of random process

the set of realizations includes infinity of possible "runs" of random process - its realizations. We will limite ourselves to finite number $\Omega$ and denote each realization as $\xi_{\omega}(t)$, or $\xi_{\omega}[n]$. In case any parameters are estimated on this set, we will speak about ensemble estimates.

Example: the random signal is recording of water flowing through the water tube in my flat. 1068 realizations, each of 20 ms , were recorded. For the demonstration of continuous random signals, we will imagine this set as $\xi_{\omega}(t)$, for discrete random signals as $\xi_{\omega}[n]$.
$\xi_{\omega}(t)$ for $\omega=1,200,500,1000$




$$
\xi_{\omega}[n] \text { for } \omega=1,200,500,1000
$$





## Distribution function

is defined for one random variable: the random process for given time $t$ or $n$ is such a random variable. Definition:

$$
\begin{aligned}
& F(x, t)=\mathcal{P}\{\xi(t)<x\} \\
& F(x, n)=\mathcal{P}\{\xi[n]<x\}
\end{aligned}
$$

where $\mathcal{P}\{\xi(t)<x\}$ or $\mathcal{P}\{\xi[n]<x\}$ is the probability that random variable in given time will be smaller than $x$. Note, that $x$ is nothing random, it is an auxiliary variable.

Ensemble estimation of distribution function: we will fix ourselves in given time $t$ or $n$, and take $\Omega$ realizations. For given $x$ we estimate:

$$
\begin{aligned}
\hat{F}(x, t) & =\frac{\sum_{\omega=1}^{\Omega} 1 \text { if } \xi_{\omega}(t)<x, 0 \text { else }}{\Omega} \\
\hat{F}(x, n) & =\frac{\sum_{\omega=1}^{\Omega} 1 \text { if } \xi_{\omega}[n]<x, \quad 0 \text { else }}{\Omega}
\end{aligned}
$$




## Probability density function

is again defined for one random variable (random process for a given time $t$ or $n$ is such a random variable). Definition:

$$
\begin{aligned}
p(x, t) & =\frac{\delta F(x, t)}{\delta x} \\
p(x, n) & =\frac{\delta F(x, n)}{\delta x}
\end{aligned}
$$

Ensemble estimation of Probability density function: The function can be obtained by numeric derivation from estimated $\hat{F}(x, t)$ or $\hat{F}(x, n)$ or it can be estimated by a histogram:

- choose given $t$ or $n$
- CHoose $L$ values $x$ from $x_{\text {min }}$ till $x_{\text {max }}$, with regular step $\Delta=\frac{x_{\max }-x_{\min }}{L-1}$ :

$$
\begin{gathered}
x_{1}=x_{\min }, \quad x_{2}=x_{\min }+\Delta, \quad x_{3}=x_{\min }+2 \Delta \ldots \\
\ldots \quad x_{L-1}=x_{\min }+(L-2) \Delta, \quad x_{L}=x_{\min }+(L-1) \Delta=x_{\max }
\end{gathered}
$$

In such a way, we'll obtain $L$ "cages" with width $\Delta$, for $x_{i}$, given cage $c h_{i}$ is from
$x_{i}-\frac{\Delta}{2}$ till $x_{i}+\frac{\Delta}{2}$. The left edge of the
left-most cage (1) will be stretched till $-\infty$, the right edge of the right-most $(L)$ till $+\infty$.


- the estimate for $x_{i}$ is computed

$$
\begin{aligned}
\hat{p}\left(x_{i}, t\right) & =\frac{\sum_{\omega=1}^{\Omega} 1 \text { if } \xi_{\omega}(t) \in c h_{i}, \quad 0 \text { else }}{\Omega \Delta} \\
\hat{p}\left(x_{i}, n\right) & =\frac{\sum_{\omega=1}^{\Omega} 1 \text { if } \xi_{\omega}[n] \in c h_{i}, \quad 0 \text { else }}{\Omega \Delta}
\end{aligned}
$$










## $F(x, t), p(x, t)$ and probabilities

our task is to determine the probability, that the value of random process in time $t$ or $n$ is in the interval $[a, b]$. We can estimate this in two ways (shown only for continuous time):

- from distribution function, having in mid its definition: $F(x, t)=\mathcal{P}\{\xi(t)<x\}$, then

$$
\mathcal{P}\{a<\xi(t)<b\}=F(b, t)-F(a, t)
$$

- from probability density function. Densities can be integrated:

$$
\mathcal{P}\{a<\xi(t)<b\}=\int_{a}^{b} p(x, t) d x
$$




Properties of distribution function and probability density function:

- the values of random process will hardly be smaller than $-\infty$, therefore $F(-\infty, t)=\mathcal{P}\{\xi(t)<-\infty\}=0$.
- the values of random process will all be smaller than $\infty$, therefore $F(+\infty, t)=\mathcal{P}\{\xi(t)<+\infty\}=1$.
- probability density function is derivation of distribution function, the inverse sense is given by an integral:

$$
F(x, t)=\int_{-\infty}^{x} p(g, t) d g
$$

as $F(+\infty, t)=1, p(x, t)$ must obey:

$$
\int_{-\infty}^{+\infty} p(x, t) d x=1
$$

- the value of probability densifty function for given $x$ is not a probability !!!
beer keg is being drunk from $x=18.00$ till 22.00. We'll define function $p(x)$ as immediate beer consumption (drinking function) and $F(x)$ as function of drunk beer (attention, $x$ is time in this example):



The behavior is simnilar to PDF and distribution functions, just replace 1 with " 1 keg":

- $F(x)$ is zero in time $-\infty$ (it is zero till 18.00), as there was no beer.
- $F(x)$ is 1 keg in time $+\infty$ (actually already at 22.00), as all beer was drunk and
there'll be no more.
- amount of drunk beer at time $x: F(x)=\int_{-\infty}^{x} p(g) d g$.
- total amount of drunk beer: $F(+\infty)=\int_{-\infty}^{+\infty} p(x) d x=1 \mathrm{keg}$.
- amount of beer drunk from time $x_{1}$ till $x_{2}$ can be computed in as difference of 2 points on $F(x)$ or by integration of $p(x)$.
- one value of $p(x)$ (for example. $p(19.00)$ ) can NOT be called "amount of drunk beer"
- noone can drink anything in infinitely short time.


## Moments

on contrary to functions, moments are values, characterizing random process in time $t$ or $n$ :
mean value or "Expectation", is the 1st moment:

$$
\begin{aligned}
& a(t)=E\{\xi(t)\}=\int_{-\infty}^{+\infty} x p(x, t) d x \\
& a[n]=E\{\xi[n]\}=\int_{-\infty}^{+\infty} x p(x, n) d x
\end{aligned}
$$

ensemble estimate of mean value for each time $t$ or $n$ the estimate is simply the mean value of samples over all realizations:

$$
\begin{aligned}
& \hat{a}(t)=\frac{1}{\Omega} \sum_{\omega=1}^{\Omega} \xi_{\omega}(t) \\
& \hat{a}[n]=\frac{1}{\Omega} \sum_{\omega=1}^{\Omega} \xi_{\omega}[n]
\end{aligned}
$$

For our signals:


variance (dispersion), standard deviation

$$
\begin{aligned}
& D(t)=E\left\{[\xi(t)-a(t)]^{2}\right\}=\int_{-\infty}^{+\infty}[x-a(t)]^{2} p(x, t) d x \\
& D[n]=E\left\{[\xi[n]-a[n]]^{2}\right\}=\int_{-\infty}^{+\infty}[x-a[n]]^{2} p(x, n) d x
\end{aligned}
$$

std - standard deviation is the square root of variance:

$$
\sigma(t)=\sqrt{D(t)} \quad \sigma[n]=\sqrt{D[n]}
$$

ensemble estimate of variance and std: for each $t$ or $n$ :

$$
\begin{array}{ll}
\hat{D}(t)=\frac{1}{\Omega} \sum_{\omega=1}^{\Omega}\left[\xi_{\omega}(t)-\hat{a}(t)\right]^{2}, & \hat{\sigma}(t)=\sqrt{\hat{D}(t)} \\
\hat{D}[n]=\frac{1}{\Omega} \sum_{\omega=1}^{\Omega}\left[\xi_{\omega}[n]-\hat{a}[n]\right]^{2}, & \hat{\sigma}[n]=\sqrt{\hat{D}[n]}
\end{array}
$$

For our signals:



## Correlation function

is quantifying the similarity between values of random process in times $t_{1}$ (or $n_{1}$ ) and $t_{2}$ (or $n_{2}$ ):

$$
\begin{aligned}
R\left(t_{1}, t_{2}\right) & =\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x_{1} x_{2} p\left(x_{1}, x_{2}, t_{1}, t_{2}\right) d x_{1} d x_{2} \\
R\left(n_{1}, n_{2}\right) & =\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x_{1} x_{2} p\left(x_{1}, x_{2}, n_{1}, n_{2}\right) d x_{1} d x_{2}
\end{aligned}
$$

where $p\left(x_{1}, x_{2}, t_{1}, t_{2}\right)$, resp. $p\left(x_{1}, x_{2}, n_{1}, n_{2}\right)$ is two-dimensional probability density function between times $t_{1}$ and $t_{2}$, resp. $n_{1}$ and $n_{2}$. Theoretically, it can be computed from two-dimensional distribution function:

$$
\begin{aligned}
& F\left(x_{1}, x_{2}, t_{1}, t_{2}\right)=\mathcal{P}\left\{\xi\left(t_{1}\right)<x_{1} \quad \text { a } \xi\left(t_{2}\right)<x_{2}\right\}, \\
& F\left(x_{1}, x_{2}, n_{1}, n_{2}\right)=\mathcal{P}\left\{\xi\left[n_{1}\right]<x_{1} \text { a } \xi\left[n_{2}\right]<x_{2}\right\}
\end{aligned}
$$

by deriving along $x_{1}$ and $x_{2}$ :

$$
\begin{aligned}
p\left(x_{1}, x_{2}, t_{1}, t_{2}\right) & =\frac{\delta^{2} F\left(x_{1}, x_{2}, t_{1}, t_{2}\right)}{\delta x_{1} \delta x_{2}} \\
p\left(x_{1}, x_{2}, n_{1}, n_{2}\right) & =\frac{\delta^{2} F\left(x_{1}, x_{2}, n_{1}, n_{2}\right)}{\delta x_{1} \delta x_{2}}
\end{aligned}
$$

We'll rather be interested in ensemble estimation using 2D-histogram:

- similarly as for 1D histogram, define "cages", in forms of squares: $c h_{i j}$ is for the first dimension from $x_{i}-\frac{\Delta}{2}$ till $x_{i}+\frac{\Delta}{2}$ and for the second dimension from $x_{j}-\frac{\Delta}{2}$ till $x_{j}+\frac{\Delta}{2}$.

- the value of 2D-histogramu for cage $c h_{i j}$ with values $x_{i}$ and $x_{j}$ will be:

$$
\begin{aligned}
\hat{p}\left(x_{i}, x_{j}, t_{1}, t_{2}\right) & =\frac{\sum_{\omega=1}^{\Omega} 1 \text { if } \xi_{\omega}\left(t_{1}\right), \xi_{\omega}\left(t_{2}\right) \in c h_{i j}, 0 \text { else }}{\Omega \Delta^{2}} \\
\hat{p}\left(x_{i}, x_{j}, n_{1}, n_{2}\right) & =\frac{\sum_{\omega=1}^{\Omega} 1 \text { if } \xi_{\omega}\left[n_{1}\right], \xi_{\omega}\left[n_{2}\right] \in c h_{i j}, \quad 0 \text { else }}{\Omega \Delta^{2}}
\end{aligned}
$$

For our signals (only discrete-time): $n_{1}=0, n_{2}=0,1,5,11$.


For $n_{1}=0, n_{2}=0,1,5,11$ the following coefficients were obtain after numeric integration:

- $R(0,0)=0.0188$ : the process is very similar to itself (the same point). similar...
- $R(0,1)=0.0151$ : if shifted to neighboring sample, still very similar to time $n_{1}=0$.
- $R(0,5)=0.0030$ : times $n_{1}=0$ and $n_{2}=5$ are not similar at all.
- $R(0,11)=-0.0133$ : in times $n_{1}=0$ and $n_{2}=11$ the process is similar to itself, but with inverse sign! It is probable, that if the value $\xi\left[n_{1}\right]$ is positive, $\xi\left[n_{2}\right]$ will be negative and vice versa.

Correlation function for



## STATIONARITY OF RANDOM PROCESS

simply said, the behavior of random process is not changing with the time. The values do not depend on time $t$ or $n$. Correlation function does not depend on the precise values of $t_{1}, t_{2}$ or $n_{1}, n_{2}$, but only on their difference: $\tau=t_{2}-t_{1}, k=n_{2}-n_{1}$. For stationary continuous-time signal:

$$
\begin{gathered}
F(x, t) \rightarrow F(x) \quad p(x, t) \rightarrow p(x) \\
a(t) \rightarrow a \quad D(t) \rightarrow D \quad \sigma(t) \rightarrow \sigma \\
p\left(x_{1}, x_{2}, t_{1}, t_{2}\right) \rightarrow p\left(x_{1}, x_{2}, \tau\right) \quad R\left(t_{1}, t_{2}\right) \rightarrow R(\tau)
\end{gathered}
$$

Similarly for discrete time:

$$
\begin{gathered}
F(x, n) \rightarrow F(x) \quad p(x, n) \rightarrow p(x) \\
a[n] \rightarrow a \quad D[n] \rightarrow D \\
p\left(x_{1}, x_{2}, n_{1}, n_{2}\right) \rightarrow p(n] \rightarrow \sigma \\
\left.x_{1}, x_{2}, k\right) \quad R\left(n_{1}, n_{2}\right) \rightarrow R(k)
\end{gathered}
$$

The "water"-example signal was obviously stationary, as:

- the mean value was similar for all times. In case we had more realizations, it would be even more similar.
- dtto for standard deviation.
- Correlation function for $n_{1}=0, n_{2}=n_{1}+k$ and for $n_{1}=100, n_{2}=n_{1}+k$ was similar.

Stationary vs. non-stationary signal:



## ERGODICITY OF RANDOM PROCESS

all parameters can be estimated from 1 realization:

Example of stationary and ergodic and of stationary but non-ergodic random process.



All ensemble estimates can be replaced by temporal estimates, we dispose of interval of lenght $T$ (continuous time) or of $N$ samples (discrete time). The only realization we have will be simply denoted $x(t)$, resp. $x[n]$ :

- histograms can be applied to estimate distribution function and PDF:
- mean value, variance, std:

$$
\begin{array}{cc}
\hat{a}=\frac{1}{T} \int_{0}^{T} x(t) d t & \hat{D}=\frac{1}{T} \int_{0}^{T}[x(t)-\hat{a}]^{2} d t \\
\hat{a}=\frac{1}{N} \sum_{n=0}^{N-1} x[n] & \hat{D}=\frac{1}{N} \sum_{n=0}^{N-1}[x[n]-\hat{D}]^{2}
\end{array} \quad \hat{\sigma}=\sqrt{\hat{D}}
$$

- correlation function:

$$
\begin{aligned}
\hat{R}(\tau) & =\frac{1}{T} \int_{0}^{T} x(t) x(t+\tau) d t \\
\hat{R}(k) & =\frac{1}{N} \sum_{n=0}^{N-1} x[n] x[n+k]
\end{aligned}
$$

